



Research article

Analysis of a stochastic SIB cholera model with saturation recovery rate and Ornstein-Uhlenbeck process

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Abstract: In this paper, a stochastic SIB(Susceptible-Infected-Vibrios) cholera model with saturation recovery rate and Ornstein-Uhlenbeck process is investigated. It is proved that there is a unique global solution for any initial value of the model. Furthermore, the sufficient criterion of the stationary distribution of the model is obtained by constructing a suitable Lyapunov function, and the expression of probability density function is calculated by the same condition. The correctness of the theoretical results is verified by numerical simulation, and the specific expression of the marginal probability density function is obtained.

Keywords: stochastic SIB cholera model; saturation recovery rate; Ornstein-Uhlenbeck process; stationary distribution; density function

1. Introduction

Cholera is an acute intestinal infectious disease caused by *Vibrio cholerae*, which is mainly transmitted through unclean water and food (see [1]). According to news reports, more than 20 countries and regions (southwestern Cameroon, Afghanistan, the Republic of Mozambique, northern Iraq, Haiti, Malawi, etc.) have experienced or are experiencing cholera epidemics in the past two years. Therefore, many scholars have studied the cholera transmission model, such as differential dynamics system [2–5], age-structured transmission model [6], reaction-convection-diffusion equations [7], generalized fractional model [8].

In [9], the scholars investigated the following SIB cholera model:

$$\begin{aligned} S'(t) &= A - \beta \frac{SB}{K+B} - \mu_1 S + \gamma I + c \frac{I}{b+I}, \\ I'(t) &= \beta \frac{SB}{K+B} - (\mu_1 + \gamma + \alpha) I - c \frac{I}{b+I}, \\ B'(t) &= \eta I - \mu_2 B. \end{aligned} \quad (1.1)$$

where $S(t)$ and $I(t)$ denote the numbers of susceptible individuals and infected individuals at time t , respectively. $B(t)$ denotes the concentration of vibrios in contaminated water at time t . Besides, The parameter A is the recruitment rate, and parameter μ_1 is the natural human death rate. The parameters β is the transmission coefficients of environment-to-human pathways. The parameter K is the concentration of vibrios in contaminated water that yields 50% chance of catching cholera. The parameter γ is the recovery rate of infected individuals. c is the maximum recovery per unit of time, and b is the infected size at 50% saturation. α is the disease-induced human death rate. The parameter η is the contribution rate of each infected individual to the concentration of vibrios, and μ_2 is the net death rate of vibrios. All parameters are usually also assumed to be nonnegative. For system (1.1), the basic reproduction number is defined by

$$R_0 = \frac{S_0 b \beta \eta}{K \mu_2 (r b + b \mu_1 + b \alpha + c)},$$

where $S_0 = \frac{A}{\mu_1}$. In [9], the scholars obtained that if $R_0 < 1$, model has only a disease-free equilibrium $E_0 = (S_0, 0, 0)$ that is locally asymptotically stable; if $R_0 > 1$, system (1.1) has a unique endemic equilibrium $E^+ = (S^+, I^+, B^+)$ that is locally asymptotically stable.

In the real environment, the spread of cholera is disturbed by random factors, so many scholars have proposed a kind of stochastic differential equation cholera spread model with random perturbation, such as [10, 11], stochastic cholera model between communities linked by migration [12], stochastic model with Lévy process [13], stochastic model under regime switching [14]. At present, the O-U process is popular in various random interference channels (see [15–18]). Therefore, in order to reveal the impact of environmental noise on the transmission rate, we assume that it is a random variable and satisfies the following form [19–22]:

$$d\gamma = \lambda_1(\bar{\gamma} - \gamma(t))dt + \sigma_1 dB_1(t), \quad dc = \lambda_2(\bar{c} - c(t))dt + \sigma_2 dB_2(t).$$

where $\bar{\gamma}, \bar{c}$ are measure the long-time mean levels of the infection rates γ, c ; $\lambda_i (i = 1, 2)$ are the speeds of reversion. $B_i(t)$ are independent standard Brownian motion parameters defined on a complete probability space (Ω, \mathcal{F}, P) , and parameter $\sigma_i > 0 (i = 1, 2)$ represents the intensity of $B_i(t)$. All parameters are usually also assumed to be nonnegative. In order to discuss the need for positivity of the stochastic model, variable $\max\{\gamma(t), 0\}, \max\{c(t), 0\}$ is used instead of variable $\gamma(t), c(t)$ in [19]. Then, we

investigated the following stochastic SIB model with O-U process:

$$\begin{aligned}
 S'(t) &= A - \beta \frac{SB}{K+B} - \mu_1 S + \max\{\gamma(t), 0\}I + \max\{c(t), 0\} \frac{I}{b+I}, \\
 I'(t) &= \beta \frac{SB}{K+B} - (\mu_1 + \max\{\gamma(t), 0\} + \alpha)I - \max\{c(t), 0\} \frac{I}{b+I}, \\
 B'(t) &= \eta I - \mu_2 B, \\
 d\gamma &= \lambda_1(\bar{\gamma} - \gamma(t))dt + \sigma_1 dB_1(t), \\
 dc &= \lambda_2(\bar{c} - c(t))dt + \sigma_2 dB_2(t).
 \end{aligned} \tag{1.2}$$

Throughout this paper, we define $\mathbb{R}_+^3 = \{(x_1, x_2, x_3) : x_i > 0, i = 1, 2, 3\}$. For an integrable function $f(t)$ defined on $[0, \infty)$, define $\langle f(t) \rangle = \frac{1}{t} \int_0^t f(s)ds$. For the numbers a and b , we define $a \vee b = \max\{a, b\}$, $a \wedge b = \min\{a, b\}$.

In the next section, we verify the existence and uniqueness of a global solution of the model (1.2) with any initial value. In Section 3, the criterion on the ergodicity and existence of unique stationary distribution for any solution of model (1.2) is stated and proved. In Section 4, We obtain the probability density function of model (1.2) around the positive equilibrium point E^* . In Section 5, the numerical examples are carried out to illustrate the main theoretical results.

2. Existence and uniqueness of a global solution

In this section, we will discuss the existence and uniqueness of a global solution for model (1.2).

Theorem 2.1. *For any initial value $(S(0), I(0), B(0), \gamma(0), c(0)) \in \mathbb{R}_+^3 \times \mathbb{R}^2$, model (1.2) has a unique global solution $(S(t), I(t), B(t), \gamma(t), c(t))$. That is, solution $(S(t), I(t), B(t), \gamma(t), c(t))$ is defined for all $t \geq 0$ and remains in $\mathbb{R}_+^3 \times \mathbb{R}^2$ with probability one.*

Proof. Since the coefficients of model (1.2) satisfy the local Lipschitz conditions, for any initial value $(S(0), I(0), B(0), \gamma(0), c(0)) \in \mathbb{R}_+^3 \times \mathbb{R}^2$, there exists a unique local solution $(S(t), I(t), B(t), \gamma(t), c(t))$ on $t \in [0, \tau_e)$, where τ_e denotes the explosion time. To show this solution is global, we only need to prove that $\tau_e = \infty$ a.s.. To this end, let $k_0 \geq 1$ be sufficiently large such that $S(0), I(0), B(0), e^{\gamma(0)}$ and $e^{c(0)}$ all lie within the interval $[\frac{1}{k_0}, k_0]$. For each integer $k \geq k_0$, define the stopping time

$$\tau_k = \inf\{t \in [0, \tau_e) : \min\{S(t), I(t), B(t), e^{\gamma(t)}, e^{c(t)}\} \leq \frac{1}{k} \text{ or } \max\{S(t), I(t), B(t), e^{\gamma(t)}, e^{c(t)}\} \geq k\},$$

where throughout this paper, we set $\inf \emptyset = \infty$ (as usual \emptyset represents the empty set). Clearly, τ_k is increasing as $k \rightarrow \infty$. Let $\tau_\infty = \lim_{k \rightarrow \infty} \tau_k$, whence $\tau_\infty \leq \tau_e$ a.s. If $\tau_\infty = \infty$ a.s. is not false, then $\tau_e = \infty$ a.s. and $(S(t), I(t), B(t), \gamma(t), c(t)) \in \mathbb{R}_+^3 \times \mathbb{R}^2$ a.s. for all $t > 0$. That is to say, if we want to finish the proof, we only need to show $\tau_\infty = \infty$ a.s. If this assertion is not true, then there is a pair of constants $T > 0$ and $\varepsilon \in (0, 1)$ such that

$$p\{\tau_\infty \leq T\} > \varepsilon.$$

Consequently, there exists an integer $k_1 \geq k_0$ such that

$$p\{\tau_k \leq T\} \geq \varepsilon \quad \text{for all } k \geq k_1. \tag{2.1}$$

Define a Lyapunov function

$$V = S - 1 - \ln S + I - 1 - \ln I + B - 1 - \ln B + \frac{1}{2}(\gamma^2 + c^2).$$

By calculating, we have

$$\begin{aligned} LV &= A - \mu_1(S + I) - \alpha I + \eta I - \mu_2 B - \frac{1}{B}[\eta I - \mu_2 B] \\ &\quad - \frac{1}{I}[\beta \frac{SB}{K+B} - (\mu_1 + \max\{\gamma(t), 0\} + \alpha)I - \max\{c(t), 0\} \frac{I}{b+I}] \\ &\quad - \frac{1}{S}[A - \beta \frac{SB}{K+B} - \mu_1 S + \max\{\gamma(t), 0\}I + \max\{c(t), 0\} \frac{I}{b+I}] \\ &\quad + \frac{1}{2}(\sigma_1^2 + \sigma_2^2) + \lambda_1 \gamma(\bar{\gamma} - \gamma) + \lambda_2 c(\bar{c} - c) \\ &\leq A + \eta I + \mu_2 + \mu_1 + |\gamma| + \alpha + |c| \frac{1}{b} + \frac{1}{2}(\sigma_1^2 + \sigma_2^2) \\ &\quad + \beta \frac{B}{K+B} + \mu_1 + \lambda_1 \gamma(\bar{\gamma} - \gamma) + \lambda_2 c(\bar{c} - c). \end{aligned} \quad (2.2)$$

By model (1.2), we have

$$\frac{d(S + I)}{dt} = A - \mu_1(S + I) - \alpha I \leq A - \mu_1(S + I). \quad (2.3)$$

This implies that

$$S + I \leq \begin{cases} S(0) + I(0), & \text{if } S(0) + I(0) \geq \frac{A}{\mu_1}, \\ \frac{A}{\mu_1}, & \text{if } S(0) + I(0) < \frac{A}{\mu_1}, \end{cases} \leq \tilde{N},$$

where $\tilde{N} = \max\{\frac{A}{\mu_1}, S(0) + I(0)\}$.

$$B'(t) = \eta I - \mu_2 B \leq \eta \tilde{N} - \mu_2 B,$$

which implies that

$$B \leq \begin{cases} B(0), & \text{if } B(0) \geq \frac{A\eta}{\mu_1\mu_2}, \\ \frac{A\eta}{\mu_1\mu_2}, & \text{if } B(0) < \frac{A\eta}{\mu_1\mu_2}, \end{cases} \leq \bar{N},$$

where $\bar{N} = \max\{\frac{A\eta}{\mu_1\mu_2}, B(0)\}$. Then,

$$\begin{aligned} LV &\leq A + \eta \tilde{N} + \mu_2 + 2\mu_1 + \alpha + \beta \frac{\tilde{N}}{K + \bar{N}} + \sup_{\gamma \in \mathbb{R}}\{|\gamma| + \lambda_1 \gamma(\bar{\gamma} - \gamma)\} \\ &\quad + \frac{1}{2}(\sigma_1^2 + \sigma_2^2) + \sup_{c \in \mathbb{R}}\{|c| \frac{1}{b} + \lambda_2 c(\bar{c} - c)\} := \tilde{K}. \end{aligned} \quad (2.4)$$

Therefore, integrating both sides of (2.4) from 0 to $\tau_k \wedge T = \min\{\tau_k, T\}$ for any $k \geq k_1$ and then taking the expectations result in

$$\begin{aligned} & \mathbb{E}V(S(\tau_k \wedge T), I(\tau_k \wedge T), B(\tau_k \wedge T), \gamma(\tau_k \wedge T), c(\tau_k \wedge T)) \\ & \leq V(S(0), I(0), B(0), \gamma(0), c(0)) + K\mathbb{E}(\tau_k \wedge T). \end{aligned}$$

Thus,

$$\begin{aligned} & \mathbb{E}V(S(\tau_k \wedge T), I(\tau_k \wedge T), B(\tau_k \wedge T), \gamma(\tau_k \wedge T), c(\tau_k \wedge T)) \\ & \leq V(S(0), I(0), B(0), \gamma(0), c(0)) + KT. \end{aligned} \quad (2.5)$$

Set $\Omega_k = \{\tau_k \leq T\}$ for $k \geq k_1$, and in view of (2.1), we get $P(\Omega_k) \geq \varepsilon$. Notice that for every $\omega \in \Omega_k$, there exists $S(\tau_k, \omega), I(\tau_k, \omega), B(\tau_k, \omega), \gamma(\tau_k, \omega)$ or $c(\tau_k, \omega)$, which equals either k or $\frac{1}{k}$. Therefore, $V(S(\tau_k, \omega), I(\tau_k, \omega), B(\tau_k, \omega), \gamma(\tau_k, \omega), c(\tau_k, \omega))$ is no less than either

$$(k - 1 - \ln k) \wedge \frac{\ln^2 k}{2} \quad \text{or} \quad \left(\frac{1}{k} - 1 + \ln k\right) \wedge \frac{\ln^2 k}{2}.$$

Thereby, we can obtain

$$V(S(\tau_k, \omega), I(\tau_k, \omega), B(\tau_k, \omega), \gamma(\tau_k, \omega), c(\tau_k, \omega)) \geq [k - 1 - \ln k] \wedge \left[\frac{1}{k} - 1 + \ln k\right] \wedge \frac{\ln^2 k}{2}.$$

By (2.5), it follows that

$$\begin{aligned} & V(S(0), I(0), B(0), \gamma(0), c(0)) + KT \\ & \geq \mathbb{E}[I_{\Omega_k}(\omega)V(S(\tau_k, \omega), I(\tau_k, \omega), B(\tau_k, \omega), \gamma(\tau_k, \omega), c(\tau_k, \omega))] \\ & \geq \varepsilon[k - 1 - \ln k] \wedge \left[\frac{1}{k} - 1 + \ln k\right] \wedge \frac{\ln^2 k}{2}, \end{aligned}$$

where I_{Ω_k} denotes the indicator function of Ω_k . Letting $k \rightarrow \infty$, then one can get that

$$\infty > V(S(0), I(0), B(0), \gamma(0), c(0)) + KT = \infty$$

which is a contradiction, and then we derive $\tau_\infty = \infty$. The proof is complete. \square

Define the set Γ as follows:

$$\Gamma = \{(S, I, B, \gamma, c) \in \mathbb{R}_+^3 \times \mathbb{R}^2 : S + I \leq \frac{A}{\mu_1}, B \leq \frac{A\eta}{\mu_1\mu_2}\}.$$

Corollary 2.2. For any initial value $x(0) = (S(0), I(0), B(0), \gamma(0), c(0)) \in \mathbb{R}_+^3 \times \mathbb{R}^2$ the global solution $x(t) = (S(t), I(t), B(t), \gamma(t), c(t))$ of model (1.2) ultimately enters into region Γ with probability one as $t \rightarrow \infty$, and when $x(0) \in \Gamma$, then $x(t) \in \Gamma$ with probability one for all $t \geq 0$.

3. Stationary distribution

In this section, we study the existence of the stationary distribution of model (1.2). Define

$$R_0^s = \frac{Ab\beta\eta}{K\mu_1\mu_2[b(\mu_1 + \bar{\gamma} + \alpha) + \bar{c} + \frac{1}{2\sqrt{\pi}}(\frac{b\sigma_1}{\sqrt{\lambda_1}} + \frac{\sigma_2}{\sqrt{\lambda_2}})]}.$$

Theorem 3.1. Assume that $R_0^s > 1$, and then model (1.2) has at least one stationary distribution and ergodic property.

Proof. Define the Lyapunov function as follows:

$$V(S, I, B, \gamma, c) = MV_1(S, I, B) + V_2(S, I, B, \gamma, c),$$

where

$$V_1 = -\ln I - c_1 \ln S - c_2 \ln B + c_3 B + c_1 \frac{\bar{\beta}_1}{K\mu_2} B,$$

$$V_2 = -\ln S - \ln B - \ln\left(\frac{A}{\mu_1} - S - I\right) + \frac{c^2 + \gamma^2}{2}.$$

By calculating, we have

$$\begin{aligned} LV_1 &= -\frac{1}{I} \left[\beta \frac{SB}{K+B} - (\mu_1 + \max\{\gamma(t), 0\} + \alpha)I - \max\{c(t), 0\} \frac{I}{b+I} \right] \\ &\quad - c_1 \frac{1}{S} \left[A - \beta \frac{SB}{K+B} - \mu_1 S + \max\{\gamma(t), 0\}I + \max\{c(t), 0\} \frac{I}{b+I} \right] \\ &\quad + c_1 \frac{\beta}{K\mu_2} [\eta I - \mu_2 B] - c_2 \frac{1}{B} [\eta I - \mu_2 B] + c_3 [\eta I - \mu_2 B] \\ &\leq -\beta \frac{SB}{I(K+B)} - c_2 \eta \frac{I}{B} - c_1 \frac{A}{S} - c_3 \mu_2 (B+K) + c_1 \mu_1 + c_2 \mu_2 + c_3 \mu_2 K + c_3 \eta I \\ &\quad + (\mu_1 + \max\{\gamma(t), 0\} + \alpha) + \frac{\max\{c(t), 0\}}{b+I} + c_1 \beta \frac{B}{K+B} + c_1 \frac{\beta}{K\mu_2} [\eta I - \mu_2 B] \\ &\leq -4 \sqrt{\beta c_2 \eta c_1 A c_3 \mu_2} + c_1 \mu_1 + c_2 \mu_2 + c_3 \mu_2 K + (\mu_1 + \bar{\gamma} + \alpha) + \frac{\bar{c}}{b} \\ &\quad + (y_1(t) \vee 0) + \frac{y_2(t) \vee 0}{b} + c_1 \frac{\beta}{K\mu_2} \eta I + c_3 \eta I \end{aligned}$$

where $y_1(t) = \gamma(t) - \bar{\gamma}$, $y_2(t) = c(t) - \bar{c}$ and

$$c_1 = \frac{A\beta\eta}{K\mu_1^2\mu_2}, c_2 = \frac{A\beta\eta}{K\mu_1\mu_2^2}, c_3 = \frac{A\beta\eta}{K^2\mu_1\mu_2^2}.$$

We have the following stochastic differential equation:

$$dy_i = -\lambda_i y_i(t) dt + \sigma_i dB_i(t), \quad i = 1, 2.$$

$y_i(t)$ has the ergodic property and density as follows:

$$\tilde{\pi}_i(x) = \frac{\sqrt{\lambda_i}}{\sqrt{\pi}\sigma_i} e^{-\frac{\lambda_i x^2}{\sigma_i^2}}, \quad x \in \mathbb{R}, \quad i = 1, 2.$$

$$\int_{-\infty}^{\infty} (x \vee 0) \tilde{\pi}_i(x) dx = \int_0^{\infty} x \frac{\sqrt{\lambda_i}}{\sqrt{\pi}\sigma_i} e^{-\frac{\lambda_i x^2}{\sigma_i^2}} dx = \frac{\sigma_i}{2\sqrt{\pi}\lambda_i}. \quad (3.1)$$

$$\begin{aligned}
LV_1 &\leq -\frac{A\beta\eta}{K\mu_1\mu_2} + (\mu_1 + \bar{\gamma} + \alpha + \frac{\bar{c}}{b}) + \frac{1}{2\sqrt{\pi}}(\frac{\sigma_1}{\sqrt{\lambda_1}} + \frac{\sigma_2}{b\sqrt{\lambda_2}}) + c_1\frac{\beta\eta}{K\mu_2}I + c_3\eta I \\
&\quad + ((y_1(t) \vee 0) - \int_0^\infty x\tilde{\pi}_1(x)dx) + \frac{1}{b}((y_2(t) \vee 0) - \int_0^\infty x\tilde{\pi}_2(x)dx) \\
&= -[\mu_1 + \bar{\gamma} + \alpha + \frac{\bar{c}}{b} + \frac{1}{2\sqrt{\pi}}(\frac{\sigma_1}{\sqrt{\lambda_1}} + \frac{\sigma_2}{b\sqrt{\lambda_2}})](R_0^s - 1) + c_1\frac{\beta\eta}{K\mu_2}I + c_3\eta I \\
&\quad + ((y_1(t) \vee 0) - \int_0^\infty x\tilde{\pi}_1(x)dx) + \frac{1}{b}((y_2(t) \vee 0) - \int_0^\infty x\tilde{\pi}_2(x)dx),
\end{aligned}$$

and

$$\begin{aligned}
LV_2 &= -\frac{1}{B}[\eta I - \mu_2 B] - \frac{1}{S}[A - \beta\frac{SB}{K+B} - \mu_1 S + \max\{\gamma(t), 0\}I + \max\{c(t), 0\}\frac{I}{b+I}] \\
&\quad + \frac{1}{\frac{A}{\mu_1} - S - I}[A - (\mu_1 + \alpha)I - \mu_1 S] + \lambda_1\gamma(\bar{\gamma} - \gamma) + \lambda_2c(\bar{c} - c) + \frac{1}{2}(\sigma_1^2 + \sigma_2^2) \\
&\leq -\frac{A}{S} - \eta\frac{I}{B} - \frac{\alpha I}{\frac{A}{\mu_1} - S - I} + \mu_2 + 2\mu_1 + \beta\frac{A\eta}{K\mu_2\mu_1 + A\eta} + \frac{1}{2}(\sigma_1^2 + \sigma_2^2) \\
&\quad + \lambda_1\gamma(\bar{\gamma} - \gamma) + \lambda_2c(\bar{c} - c).
\end{aligned}$$

Then,

$$\begin{aligned}
LV &\leq M\{-[\mu_1 + \bar{\gamma} + \alpha + \frac{\bar{c}}{b} + \frac{1}{2\sqrt{\pi}}(\frac{\sigma_1}{\sqrt{\lambda_1}} + \frac{\sigma_2}{b\sqrt{\lambda_2}})](R_0^s - 1) + c_1\frac{\beta\eta}{K\mu_2}I + c_3\eta I \\
&\quad + ((y_1(t) \vee 0) - \int_0^\infty x\tilde{\pi}_1(x)dx) + \frac{1}{b}((y_2(t) \vee 0) - \int_0^\infty x\tilde{\pi}_2(x)dx)\} \\
&\quad - \frac{A}{S} - \eta\frac{I}{B} - \frac{\alpha I}{\frac{A}{\mu_1} - S - I} + \mu_2 + 2\mu_1 + \frac{A\eta\beta}{K\mu_2\mu_1 + A\eta} + \frac{1}{2}(\sigma_1^2 + \sigma_2^2) \\
&\quad + \lambda_1\gamma(\bar{\gamma} - \gamma) + \lambda_2c(\bar{c} - c) \\
&:= F(S, I, B, \gamma, c) + M((y_1(t) \vee 0) - \int_0^\infty x\tilde{\pi}_1(x)dx) + \frac{M}{b}((y_2(t) \vee 0) - \int_0^\infty x\tilde{\pi}_2(x)dx),
\end{aligned} \tag{3.2}$$

where

$$\begin{aligned}
F(S, I, B, \gamma, c) &= -M[\mu_1 + \bar{\gamma} + \alpha + \frac{\bar{c}}{b} + \frac{1}{2\sqrt{\pi}}(\frac{\sigma_1}{\sqrt{\lambda_1}} + \frac{\sigma_2}{b\sqrt{\lambda_2}})](R_0^s - 1) + \mu_2 + 2\mu_1 \\
&\quad + (\frac{c_1\beta}{K\mu_2}I + c_3)M\eta I - \frac{\alpha I}{\frac{A}{\mu_1} - S - I} + \frac{A\eta\beta}{K\mu_2\mu_1 + A\eta} + \frac{\sigma_1^2 + \sigma_2^2}{2} - \frac{A}{S} \\
&\quad - \eta\frac{I}{B} - \frac{1}{2}\lambda_1\gamma^2 - \frac{1}{2}\lambda_2c^2 + \sup_{(\gamma, c)^T \in \mathbb{R}^2} \{\lambda_1\gamma(\bar{\gamma} - \frac{1}{2}\gamma) + \lambda_2c(\bar{c} - \frac{1}{2}c)\}.
\end{aligned}$$

Define the bounded closed set

$$U = \{(S, I, B, \gamma, c) \in \Gamma \mid S + I \leq \frac{A}{\mu_1} - \varepsilon^2, \varepsilon \leq S, \varepsilon \leq I, \varepsilon \leq B, |\gamma| \leq \frac{1}{\varepsilon}, |c| \leq \frac{1}{\varepsilon}\},$$

where ε are small enough positive constants, which will be determined later .

For convenience, we divide $\mathbb{R}_+^3 \times \mathbb{R}^2 \setminus U$ into six domains.

$$\begin{aligned} U_1 &= \{(S, I, B, \gamma, c) \in \mathbb{R}_+^3 \times \mathbb{R}^2, 0 < S < \varepsilon\}, \\ U_2 &= \{(S, I, B, \gamma, c) \in \mathbb{R}_+^3 \times \mathbb{R}^2, 0 < I < \varepsilon\}, \\ U_3 &= \{(S, I, B, \gamma, c) \in \mathbb{R}_+^3 \times \mathbb{R}^2, 0 < B < \varepsilon^2, I > \varepsilon\}, \\ U_4 &= \{(S, I, B, \gamma, c) \in \mathbb{R}_+^3 \times \mathbb{R}^2, S + I > \frac{A}{\mu_1} - \varepsilon^2, I > \varepsilon\}, \\ U_5 &= \{(S, I, B, \gamma, c) \in \mathbb{R}_+^3 \times \mathbb{R}^2, |\gamma| \geq \frac{1}{\varepsilon}\}, \\ U_6 &= \{(S, I, B, \gamma, c) \in \mathbb{R}_+^3 \times \mathbb{R}^2, |c| \geq \frac{1}{\varepsilon}\}. \end{aligned}$$

We will prove that $F(S, I, B, \gamma, c) \leq -1$ on $\mathbb{R}_+^3 \times \mathbb{R}^2 \setminus U$, which is equivalent to show it on the above seven domains.

Case 1. If $(S, I, B, \gamma, c) \in U_1$, we can obtain

$$F(S, I, B, \gamma, c) \leq -\frac{A}{S} + G_1 \leq -\frac{A}{\varepsilon} + G_1,$$

where

$$\begin{aligned} G_1 &= -M[\mu_1 + \bar{\gamma} + \alpha + \frac{\bar{c}}{b} + \frac{1}{2\sqrt{\pi}}(\frac{\sigma_1}{\sqrt{\lambda_1}} + \frac{\sigma_2}{b\sqrt{\lambda_2}})](R_0^s - 1) + (\frac{c_1\beta}{K\mu_2}I + c_3)M\eta I + 2\mu_1 \\ &\quad + \mu_2 + \frac{A\eta\beta}{K\mu_2\mu_1 + A\eta} + \frac{1}{2}(\sigma_1^2 + \sigma_2^2) + \sup_{(\gamma, c)^T \in \mathbb{R}^2} \{\lambda_1\gamma(\bar{\gamma} - \frac{1}{2}\gamma) + \lambda_2c(\bar{c} - \frac{1}{2}c)\}. \end{aligned}$$

We choose a constant $\varepsilon > 0$ small enough such that $-\frac{A}{\varepsilon} + G_1 \leq -1$, and then it follows that

$$F(S, I, B, \gamma, c) \leq -1 \quad \text{for all } (S, I, B, \gamma, c) \in U_1. \quad (3.3)$$

Case 2. If $(S, I, B, \gamma, c) \in U_2$, we can obtain

$$F(S, I, B, \gamma, c) \leq (c_1\frac{\beta}{K\mu_2} + c_3)M\eta I + G_2 \leq (c_1\frac{\beta}{K\mu_2} + c_3)M\eta\varepsilon + G_2,$$

where

$$\begin{aligned} G_2 &= -M[\mu_1 + \bar{\gamma} + \alpha + \frac{\bar{c}}{b} + \frac{1}{2\sqrt{\pi}}(\frac{\sigma_1}{\sqrt{\lambda_1}} + \frac{\sigma_2}{b\sqrt{\lambda_2}})](R_0^s - 1) + \frac{1}{2}(\sigma_1^2 + \sigma_2^2) \\ &\quad + 2\mu_1 + \mu_2 + \frac{A\eta\beta}{K\mu_2\mu_1 + A\eta} + \sup_{(\gamma, c)^T \in \mathbb{R}^2} \{\lambda_1\gamma(\bar{\gamma} - \frac{1}{2}\gamma) + \lambda_2c(\bar{c} - \frac{1}{2}c)\}. \end{aligned}$$

Choose constants $M > 0$ large enough and $\varepsilon > 0$ small enough such that $(c_1\frac{\beta}{K\mu_2} + c_3)M\eta\varepsilon + G_2 \leq -1$, and then it follows that

$$F(S, I, B, \gamma, c) \leq -1 \quad \text{for all } (S, I, B, \gamma, c) \in U_2. \quad (3.4)$$

Case 3. If $(S, I, B, \gamma, c) \in U_3$, we can obtain

$$F(S, I, B, \gamma, c) \leq -\eta \frac{I}{B} + G_1 \leq -\eta \frac{\varepsilon}{\varepsilon^2} + G_1.$$

Choose a constant $\varepsilon > 0$ small enough such that $-\frac{\eta}{\varepsilon} + G_1 \leq -1$, and then it follow that

$$F(S, I, B, \gamma, c) \leq -1 \quad \text{for all } (S, I, B, \gamma, c) \in U_3. \quad (3.5)$$

Case 4. If $(S, I, B, \gamma, c) \in U_4$, we can obtain

$$F(S, I, B, \gamma, c) \leq -\frac{\alpha I}{\frac{A}{\mu_1} - S - I} + G_1 \leq -\alpha \frac{\varepsilon}{\varepsilon^2} + G_1.$$

Choose a constant $\varepsilon > 0$ small enough such that $-\frac{\alpha}{\varepsilon} + G_1 \leq -1$, and then we have

$$F(S, I, B, \gamma, c) \leq -1 \quad \text{for all } (S, I, B, \gamma, c) \in U_4. \quad (3.6)$$

Case 5. If $(S, I, B, \gamma, c) \in U_5$, we can obtain

$$F(S, I, B, \gamma, c) \leq -\frac{1}{2}\gamma^2 + G_1 \leq -\frac{1}{2\varepsilon^2} + G_1.$$

Choose a constant $\varepsilon > 0$ small enough such that $-\frac{1}{2\varepsilon^2} + G_1 \leq -1$, and then we get

$$F(S, I, B, \gamma, c) \leq -1 \quad \text{for all } (S, I, B, \gamma, c) \in U_5. \quad (3.7)$$

Case 6. If $(S, I, B, \gamma, c) \in U_6$, we can obtain

$$F(S, I, B, \gamma, c) \leq -\frac{1}{2}c^2 + G_1 \leq -\frac{1}{2\varepsilon^2} + G_1.$$

Choose a constant $\varepsilon > 0$ small enough such that $-\frac{1}{2\varepsilon^2} + G_1 \leq -1$, and then we get

$$F(S, I, B, \gamma, c) \leq -1 \quad \text{for all } (S, I, B, \gamma, c) \in U_6. \quad (3.8)$$

Finally, from (3.3)-(3.8) we obtain

$$F(S, I, B, \gamma, c) \leq -1 \quad \text{for all } (S, I, B, \gamma, c) \in \mathbb{R}_+^3 \times \mathbb{R}^2 \setminus U. \quad (3.9)$$

Define

$$\bar{V}(S, I, B, \gamma, c) = V(S, I, B, \gamma, c) - V(S^0, I^0, B^0, \gamma^0, c^0),$$

where $(S^0, I^0, B^0, \gamma^0, c^0)$ is the point in the interior $\mathbb{R}_+^3 \times \mathbb{R}^2$, such that $V(S, I, B, \gamma, c)$ will be minimized. By (3.2), we have

$$L\bar{V} \leq F(S, I, B, \gamma, c) + M((y_1(t) \vee 0) - \int_0^\infty x\tilde{\pi}_1(x)dx) + \frac{M}{b}((y_2(t) \vee 0) - \int_0^\infty x\tilde{\pi}_2(x)dx). \quad (3.10)$$

Taking the mathematical expectation and Itô's integral of (3.10), for any initial value $(S(0), I(0), B(0), \gamma(0), c(0))$, one has

$$\begin{aligned} 0 &\leq \frac{\mathbb{E}\bar{V}(S(t), I(t), B(t), \gamma(t), c(t))}{t} \\ &\leq \frac{\mathbb{E}\bar{V}(S(0), I(0), B(0), \gamma(0), c(0))}{t} + \frac{1}{t} \int_0^t \mathbb{E}(F(S(s), I(s), B(s), \gamma(s), c(s)))ds \\ &\quad + M[\mathbb{E}(\frac{1}{t} \int_0^t (y_1 \vee 0)ds - \int_0^\infty x\tilde{\pi}_1(x)dx) + \frac{1}{b}\mathbb{E}(\frac{1}{t} \int_0^t (y_2 \vee 0)ds - \int_0^\infty x\tilde{\pi}_2(x)dx)]. \end{aligned} \quad (3.11)$$

By the ergodicity of $y_i (i = 1, 2)$ and the strong law of large numbers, one has

$$\begin{aligned} &\lim_{t \rightarrow \infty} \mathbb{E}[\frac{1}{t} \int_0^t (y_i \vee 0)ds - \int_0^\infty x\tilde{\pi}_i(x)dx] \\ &= \mathbb{E}[\int_0^\infty x\tilde{\pi}_i(x)dx] - \int_0^\infty x\tilde{\pi}_i(x)dx = 0 \text{ a.s.} \end{aligned} \quad (3.12)$$

By (3.12), we have the inferior limit of (3.11),

$$\begin{aligned} 0 &\leq \liminf_{t \rightarrow \infty} \frac{\mathbb{E}\bar{V}(S(0), I(0), B(0), \gamma(0), c(0))}{t} \\ &\quad + \liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbb{E}(F(S(s), I(s), B(s), \gamma(s), c(s)))ds \\ &= \liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbb{E}(F(S(s), I(s), B(s), \gamma(s), c(s)))ds. \end{aligned} \quad (3.13)$$

From (3.9), we have

$$\begin{aligned} &\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbb{E}(F(S(s), I(s), B(s), \gamma(s), c(s)))ds \\ &= \liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbb{E}(F(S(s), I(s), B(s), \gamma(s), c(s)))\mathbf{1}_{(S(s), I(s), B(s), \gamma(s), c(s)) \in U} ds \\ &\quad + \liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbb{E}(F(S(s), I(s), B(s), \gamma(s), c(s)))\mathbf{1}_{(S(s), I(s), B(s), \gamma(s), c(s)) \in (\mathbb{R}_+^3 \times \mathbb{R}^2 \setminus U)} ds \\ &\leq \check{F} \liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbf{1}_{(S(s), I(s), B(s), \gamma(s), c(s)) \in U} ds + \liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbf{1}_{(S(s), I(s), B(s), \gamma(s), c(s)) \in (\mathbb{R}_+^3 \times \mathbb{R}^2 \setminus U)} ds \\ &\leq -1 + (\check{F} + 1) \liminf_{t \rightarrow +\infty} \frac{1}{t} \int_0^t \mathbf{1}_{(S(s), I(s), B(s), \gamma(s), c(s)) \in U} ds. \end{aligned} \quad (3.14)$$

By (3.13) and (3.14), we have

$$\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t \mathbf{1}_{(S(s), I(s), B(s), \gamma(s), c(s)) \in U} ds \geq \frac{1}{\check{F} + 1} > 0 \text{ a.s.} \quad (3.15)$$

In view of the definition of event probability and Fatou's lemma (see [23]), the result of (3.15) is

$$\liminf_{t \rightarrow \infty} \frac{1}{t} \int_0^t P(s, (S(s), I(s), B(s), \gamma(s), c(s)), U) ds \geq \frac{1}{\check{F} + 1} > 0 \text{ a.s.} \quad (3.16)$$

where $P(t, (S, I, B, \gamma, c), U)$ is the transition probability of $(S(t), I(t), B(t), \gamma(t), c(t))$ belonging to set U . This shows that the solution $(S(t), I(t), B(t), \gamma(t), c(t))$ of model (1.2) has the Feller and ergodic property. This completes the proof. \square

4. Density function

The positive equilibrium $P^* = (S^*, I^*, B^*, \gamma^*, c^*)$ of system (1.2) satisfied the following equations:

$$\begin{cases} 0 = A - \beta \frac{S^* B^*}{K + B^*} - \mu_1 S^* + \max\{\gamma^*, 0\}I + \max\{c^*, 0\} \frac{I^*}{b + I^*}, \\ 0 = \beta \frac{S^* B^*}{K + B^*} - (\mu_1 + \max\{\gamma^*, 0\} + \alpha)I^* - \max\{c^*, 0\} \frac{I}{b + I^*}, \\ 0 = \eta I^* - \mu_2 B^*, \\ 0 = \lambda_1(\bar{\gamma} - \gamma^*), \\ 0 = \lambda_2(\bar{c} - c^*). \end{cases} \quad (4.1)$$

When $R_0 > 1$, we have

$$I^* = I^+, \quad S^* = S^+, \quad B^* = B^+, \quad \gamma^* = \bar{\gamma}, \quad c^* = \bar{c}.$$

Let $(x_1, x_2, x_3, y_1, y_2) = (S - S^*, I - I^*, B - B^*, \gamma - \gamma^*, c - c^*)$. The linearization model of system (1.2) is:

$$\begin{cases} dx_1(t) = [-a_{11}x_1 + a_{12}x_2 - a_{13}x_3 + a_{14}y_1 + a_{15}y_2]dt, \\ dx_2(t) = [a_{21}x_1 - a_{22}x_2 + a_{13}x_3 - a_{14}y_1 - a_{15}y_2]dt, \\ dx_3(t) = [\eta x_2 - \mu_2 x_3]dt, \\ dy_1(t) = -\lambda_1 y_1 dt + \sigma_1 dB_1(t), \\ dy_2(t) = -\lambda_2 y_2 dt + \sigma_2 dB_2(t), \end{cases} \quad (4.2)$$

where

$$\begin{aligned} a_{11} &= \beta \frac{B^*}{K + B^*} + \mu_1, & a_{12} &= \gamma^* + \frac{c^* b}{(b + I^*)^2}, & a_{13} &= \frac{S^* K \beta}{(K + B^*)^2}, & a_{14} &= I^*, \\ a_{15} &= \frac{I^*}{b + I^*}, & a_{21} &= \beta \frac{B^*}{K + B^*}, & a_{22} &= (\mu_1 + \gamma^* + \alpha) + \frac{c^* b}{(b + I^*)^2}. \end{aligned}$$

Define

$$D = \begin{pmatrix} -a_{11} & a_{12} & -a_{13} & a_{14} & a_{15} \\ a_{21} & -a_{22} & a_{13} & -a_{14} & -a_{15} \\ 0 & \eta & -\mu_2 & 0 & 0 \\ 0 & 0 & 0 & -\lambda_1 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_2 \end{pmatrix}, \quad Q = \begin{pmatrix} 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & \sigma_1 & 0 \\ 0 & 0 & 0 & 0 & \sigma_2 \end{pmatrix}.$$

Let $X(t) = (x_1(t), x_2(t), x_3(t), y_1(t), y_2(t))^T$, $P(t) = (0, 0, 0, 0, B_1(t), B_2(t))^T$, and then the linear system (4.2) can be written as follows:

$$dX(t) = DX(t)dt + QdP(t).$$

The characteristic equation of

$$\phi(\lambda_A) = (\lambda + \lambda_1)(\lambda + \lambda_2)(\lambda^3 - a_1 \lambda^2 - a_2 \lambda - a_3) = 0,$$

where

$$\begin{aligned} a_1 &= a_{11} + a_{22} + \mu_2 > 0, \\ a_2 &= (a_{11} + a_{22})\mu_2 - a_{13}\eta + a_{11}a_{22} - a_{21}a_{12} > 0, \\ a_3 &= \mu_2(a_{11}a_{22} - a_{21}a_{12}) + (a_{21} - a_{11})a_{13}\eta > 0, \end{aligned} \quad (4.3)$$

and

$$a_2a_1 - a_3 = (a_{11} + a_{22})[(a_{11} + a_{22})\mu_2 - a_{13}\eta + a_{11}a_{22} - a_{21}a_{12}] + \mu_2[(a_{11} + a_{22})\mu_2 - a_{13}\eta] + (a_{11} - a_{21})a_{13}\eta > 0. \quad (4.4)$$

Lemma 4.1. For the five-dimensional algebraic equation $P_0^2 + D_0\Sigma_0 + \Sigma_0D_0^T = 0$, where $P_0 = \text{diag}(1, 0, 0, 0, 0)$, Σ_0 is a symmetrical matrix,

(1)

$$D_0 = \begin{pmatrix} -a_1 & -a_2 & -a_3 & -a_4 & -a_5 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & \bar{a} & 0 \\ 0 & 0 & 0 & 0 & \tilde{a} \end{pmatrix}.$$

If D_0 is a Hurwitz matrix, that is, $a_i > 0 (i = 1, 2, 3)$, $a_2a_1 - a_3 > 0$, then Σ_0 is positive definite and has the following expression:

$$\Sigma_0 = \begin{pmatrix} \frac{a_2}{2(a_1a_2 - a_3)} & 0 & -\frac{1}{2(a_1a_2 - a_3)} & 0 & 0 \\ 0 & \frac{1}{2(a_1a_2 - a_3)} & 0 & 0 & 0 \\ -\frac{1}{2(a_1a_2 - a_3)} & 0 & \frac{a_1}{2a_3(a_1a_2 - a_3)} & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix}.$$

(2)

$$D_0 = \begin{pmatrix} -a_1 & -a_2 & -a_3 & -a_4 & -a_5 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & a \end{pmatrix}.$$

If D_0 is a Hurwitz matrix, that is, $a_i > 0 (i = 1, 2, 3, 4)$, $a_2a_1 - a_3 > 0$ and, $a_1a_2a_3 - a_3^2 - a_1^2a_4 > 0$, then Σ_0 is positive semi-definite and has the following expression:

$$\Sigma_0 = \begin{pmatrix} \frac{a_2a_3 - a_1a_4}{2(a_1a_2a_3 - a_1^2a_4 - a_3^2)} & 0 & -\frac{a_3}{2(a_1a_2a_3 - a_1^2a_4 - a_3^2)} & 0 & 0 \\ 0 & \frac{a_3}{2(a_1a_2a_3 - a_1^2a_4 - a_3^2)} & 0 & -\frac{a_1}{2(a_1a_2a_3 - a_1^2a_4 - a_3^2)} & 0 \\ -\frac{a_3}{2(a_1a_2a_3 - a_1^2a_4 - a_3^2)} & 0 & \frac{a_1}{2(a_1a_2a_3 - a_1^2a_4 - a_3^2)} & 0 & 0 \\ 0 & -\frac{a_1}{2(a_1a_2a_3 - a_1^2a_4 - a_3^2)} & 0 & \frac{a_1a_2 - a_3}{2a_4(a_1a_2a_3 - a_1^2a_4 - a_3^2)} & 0 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix}.$$

Theorem 4.2. If $R_0^S > 1$, the stationary solution $(S(t), I(t), B(t), \gamma(t), c(t))$ of model (1.2) around $(S^*, I^*, B^*, \gamma^*, c^*)^T$ has a unique normal density function, which takes the form

$$\Phi(S, I, B, \gamma, c) = (2\pi)^{-\frac{5}{2}} |\Sigma|^{-\frac{1}{2}} e^{-\frac{1}{2}(S - S^*, I - I^*, B - B^*, \gamma - \gamma^*, c - c^*)\Sigma^{-1}(S - S^*, I - I^*, B - B^*, \gamma - \gamma^*, c - c^*)^T},$$

with the positive definite matrix $\Sigma = \Sigma_1 + \Sigma_2$.

(1) $\mu_1 \neq \mu_2$

$$\Sigma_1 = (a_{14}\eta(a_{21} - a_{11} + \mu_2)\sigma_1)^2 (H_1 H_2 H_3 H_4 H_5)^{-1} \Sigma_{11} [(H_1 H_2 H_3 H_4 H_5)^{-1}]^T.$$

$$\Sigma_2 = (a_{15}\eta(a_{21} - a_{11} + \mu_2)\sigma_2)^2 (H_8 H_2 H_3 H_4 H_9)^{-1} \Sigma_{21} [(H_8 H_2 H_3 H_4 H_9)^{-1}]^T.$$

(2) $\mu_1 = \mu_2$

$$\Sigma_1 = (a_{14}\alpha\sigma_1)^2 (H_1 H_2 H_3 H_6 H_7)^{-1} \Sigma_{12} [(H_1 H_2 H_3 H_6 H_7)^{-1}]^T.$$

$$\Sigma_2 = (a_{15}\alpha\sigma_2)^2 (H_8 H_2 H_3 H_6 H_{10})^{-1} \Sigma_{22} [(H_8 H_2 H_3 H_6 H_{10})^{-1}]^T.$$

$H_i (i = 1, 2, 3, 4, 5, 6, 7, 8, 9, 10)$, $\Sigma_{ij} (j, i = 1, 2)$ are given in the following proof.

Proof. Let the symmetric matrices $\Sigma = (r_{ij})_{5 \times 5}$, where $r_{ji} = r_{ij}$. Σ is determined by the following algebraic equation:

$$Q^2 + D\Sigma + \Sigma D^T = 0. \quad (4.5)$$

Furthermore, let $\Sigma = \Sigma_1 + \Sigma_2$, $Q_1 = \text{diag}(0, 0, 0, 0, \sigma_1, 0)$, $Q_2 = \text{diag}(0, 0, 0, 0, 0, \sigma_2)$, and then equation (4.5) can be decomposed into the following two equations: $Q_i^2 + D\Sigma_i + \Sigma_i D^T = 0$ ($i = 1, 2$). We will calculate the matrices Σ_i ($i = 1, 2$) in the following two steps.

Step (1) $Q_1^2 + D\Sigma_1 + \Sigma_1 D^T = 0$. Let $D_1 = H_1 D H_1^{-1}$, where

$$H_1 = \begin{pmatrix} 0 & 0 & 0 & 1 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix}, D_1 = \begin{pmatrix} -\lambda_1 & 0 & 0 & 0 & 0 \\ -a_{14} & -a_{22} & a_{21} & a_{13} & -a_{15} \\ a_{14} & a_{12} & -a_{11} & -a_{13} & a_{15} \\ 0 & \eta & 0 & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_2 \end{pmatrix}.$$

Let $D_2 = H_2 D_1 H_2^{-1}$, where

$$H_2 = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 1 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix}, D_2 = \begin{pmatrix} -\lambda_1 & 0 & 0 & 0 & 0 \\ -a_{14} & -a_{22} - a_{21} & a_{21} & a_{13} & -a_{15} \\ 0 & a_{12} - a_{22} - (a_{21} - a_{11}) & a_{21} - a_{11} & 0 & 0 \\ 0 & \eta & 0 & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_2 \end{pmatrix}.$$

where $a_{12} - a_{22} - (a_{21} - a_{11}) = -\alpha$. Let $D_3 = H_3 D_2 H_3^{-1}$, where

$$H_3 = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & \frac{\eta}{\alpha} & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix}, D_3 = \begin{pmatrix} -\lambda_1 & 0 & 0 & 0 & 0 \\ -a_{14} & -a_{22} - a_{21} & a_{21} - \frac{\eta}{\alpha} a_{13} & a_{13} & -a_{15} \\ 0 & -\alpha & a_{21} - a_{11} & 0 & 0 \\ 0 & 0 & \frac{\eta}{\alpha} (a_{21} - a_{11} + \mu_2) & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_2 \end{pmatrix}.$$

Case 1. Let $a_{21} - a_{11} + \mu_2 \neq 0$, that is, $\mu_2 \neq \mu_1$. Let $D_4 = H_4 D_3 H_4^{-1}$, where

$$H_4 = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 \\ 0 & -\eta(a_{21} - a_{11} + \mu_2) & \frac{\eta}{\alpha} [(a_{21} - a_{11})^2 - \mu_2^2] & \mu_2^2 & 0 \\ 0 & 0 & \frac{\eta}{\alpha} (a_{21} - a_{11} + \mu_2) & -\mu_2 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix}.$$

$$D_4 = \begin{pmatrix} -\lambda_1 & 0 & 0 & 0 & 0 \\ a_{14}\eta(a_{21} - a_{11} + \mu_2) & -a_1 & -a_2 & -a_3 & a_{15}\eta(a_{21} - a_{11} + \mu_2) \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_2 \end{pmatrix}.$$

Let $D_5 = H_5 D_4 H_5^{-1}$, where

$$H_5 = \begin{pmatrix} a_{14}\eta(a_{21} - a_{11} + \mu_2) & -a_1 & -a_2 & -a_3 & a_{15}\eta(a_{21} - a_{11} + \mu_2) \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix},$$

where a_1, a_2, a_3 are given in (4.3).

$$D_5 = \begin{pmatrix} -(a_1 + \lambda_1) & -(a_1\lambda_1 + a_2) & -(a_2\lambda_1 + a_3) & -a_3\lambda_1 & -a_{15}\eta(a_{21} - a_{11} + \mu_2)(a_1 + \lambda_2) \\ 1 & 0 & 0 & 0 & a_{15}\eta(a_{21} - a_{11} + \mu_2) \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_2 \end{pmatrix}.$$

Then, by Lemma 4.1(2), for the matrix equation

$$Q_0^2 + D_5 \Sigma_{11} + \Sigma_{11} D_5^T = 0,$$

with $Q_0 = \text{diag}(1, 0, 0, 0, 0)$, we can obtain the positive definite matrix as follows:

$$\Sigma_{11} = \begin{pmatrix} \frac{(a_1 a_2 - a_3)\lambda_1^2 + a_2^2 \lambda_1 + a_2 a_3}{2H_{11}} & 0 & -\frac{a_2 \lambda_1 + a_3}{2H_{11}} & 0 & 0 \\ 0 & \frac{a_2 \lambda_1 + a_3}{2H_{11}} & 0 & -\frac{a_1 + \lambda_1}{2H_{11}} & 0 \\ -\frac{a_2 \lambda_1 + a_3}{2H_{11}} & 0 & \frac{a_1 + \lambda_1}{2H_{11}} & 0 & 0 \\ 0 & -\frac{a_1 + \lambda_1}{2H_{11}} & 0 & \frac{a_1 a_2 - a_3 + a_1 \lambda_1 (a_1 + \lambda_1)}{2a_3 \lambda_1 H_{11}} & 0 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix},$$

where $H_{11} = (a_1 a_2 - a_3)(a_3 + \lambda_1 a_2 + \lambda_1^2 a_1 + \lambda_1^3)$. The element $a_{15}\eta(a_{21} - a_{11} + \mu_2)$ in the second row and fifth column of matrix D_5 does not affect the calculation results using the method of Lemma 4.1(2).

Then, we have that $Q_1^2 + A \Sigma_1 + \Sigma_1 A^T = 0$ is equivalent to

$$(H_1 H_2 H_3 H_4 H_5) Q_1^2 (H_1 H_2 H_3 H_4 H_5)^T + D_5 (H_1 H_2 H_3 H_4 H_5) \Sigma_1 (H_1 H_2 H_3 H_4 H_5)^T + (H_1 H_2 H_3 H_4 H_5) \Sigma_1 (H_1 H_2 H_3 H_4 H_5)^T D_5^T = 0,$$

which is further equivalent to

$$Q_0^2 + (-a_{14}\eta(a_{21} - a_{11} + \mu_2)\sigma_1)^{-2} (D_5 (H_1 H_2 H_3 H_4 H_5) \Sigma_1 (H_1 H_2 H_3 H_4 H_5)^T + (H_1 H_2 H_3 H_4 H_5) \Sigma_1 (H_1 H_2 H_3 H_4 H_5)^T D_5^T) = 0.$$

Therefore, we finally have that

$$\Sigma_1 = (a_{14}\eta(a_{21} - a_{11} + \mu_2)\sigma_1)^2(H_1H_2H_3H_4H_5)^{-1}\Sigma_{11}[(H_1H_2H_3H_4H_5)^{-1}]^T.$$

Thus, Σ_1 is calculated and is positive semi-definite.

Case 2. $\mu_2 = \mu_1$. Let $D_6 = H_6D_3H_6^{-1}$, where

$$H_6 = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 \\ 0 & -\alpha & a_{21} - a_{11} & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix}, D_6 = \begin{pmatrix} -\lambda_1 & 0 & 0 & 0 & 0 \\ a_{14}\alpha & -a_{11} - a_{22} & -D^* & -a_{13}\alpha & a_{15}\alpha \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_2 \end{pmatrix},$$

where $D^* = (a_{22} + a_{21})(a_{11} - a_{21}) + a_{21}\alpha - a_{13}\eta$. Let $D_7 = H_7D_6H_7^{-1}$, where

$$H_7 = \begin{pmatrix} a_{14}\alpha & -a_{11} - a_{22} & -D^* & -a_{13}\alpha & a_{15}\alpha \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix},$$

$$D_7 = \begin{pmatrix} -(a_{11} + a_{22} + \lambda_1) & -((a_{11} + a_{22})\lambda_1 + D^*) & -D^*\lambda_1 & D_{71} & D_{72} \\ 1 & 0 & 0 & -a_{13}\alpha & a_{15}\alpha \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_2 \end{pmatrix},$$

where $D_{71} = -a_{13}\alpha(a_{11} + a_{22} + \mu_2)$, $D_{72} = -a_{15}\alpha(a_{11} + a_{22} - \lambda_2)$. Then, by Lemma 4.1(1), for the matrix equation

$$Q_0^2 + D_7\Sigma_{12} + \Sigma_{12}D_7^T = 0,$$

we can obtain the positive define matrix as follows:

$$\Sigma_{12} = \begin{pmatrix} \frac{(a_{11}+a_{22})\lambda_1+D^*}{2H_{12}} & 0 & -\frac{1}{2H_{12}} & 0 & 0 \\ 0 & \frac{1}{2H_{12}} & 0 & 0 & 0 \\ -\frac{1}{2H_{12}} & 0 & \frac{a_{11}+a_{22}+\lambda_1}{2D^*\lambda_1H_{12}} & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix}.$$

$H_{12} = (a_{11} + a_{22} + \lambda_1)a_{11} + a_{22}\lambda_1 + a_{11} + a_{22}D^*$. The elements $-a_{13}\alpha, a_{15}\alpha$ in the second row, fourth column and fifth column of matrix D_5 do not affect the calculation results using the method of Lemma 4.1(1). Then, we have that $Q_1^2 + D\Sigma_1 + \Sigma_1D^T = 0$ is equivalent to

$$(H_1H_2H_3H_6H_7)Q_1^2(H_1H_2H_3H_6H_7)^T + D_7(H_1H_2H_3H_6H_7)\Sigma_1(H_1H_2H_3H_6H_7)^T + (H_1H_2H_3H_6H_7)\Sigma_1(H_1H_2H_3H_6H_7)^T D_7^T = 0,$$

which is further equivalent to

$$Q_0^2 + (a_{14}\alpha\sigma_1)^{-2}(D_7(H_1H_2H_3H_6H_7)\Sigma_1(H_1H_2H_3H_6H_7)^T + (H_1H_2H_3H_6H_7)\Sigma_1(H_1H_2H_3H_6H_7)^T D_7^T) = 0.$$

Therefore, we finally have that

$$\Sigma_1 = (a_{14}\alpha\sigma_1)^2(H_1H_2H_3H_6H_7)^{-1}\Sigma_{12}[(H_1H_2H_3H_6H_7)^{-1}]^T.$$

Thus, Σ_1 is calculated and is positive semi-definite.

Step (2) $Q_2^2 + D\Sigma_2 + \Sigma_2D^T = 0$. Let $D_8 = H_8DH_8^{-1}$, $D_9 = H_2D_8H_2^{-1}$, $D_{10} = H_3D_9H_3^{-1}$, where H_2, H_3 are given in **step (1)**, and

$$H_8 = \begin{pmatrix} 0 & 0 & 0 & 0 & 1 \\ 0 & 1 & 0 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \end{pmatrix}, D_8 = \begin{pmatrix} -\lambda_2 & 0 & 0 & 0 & 0 \\ -a_{15} & -a_{22} & a_{21} & a_{13} & -a_{14} \\ a_{15} & a_{12} & -a_{11} & -a_{13} & a_{14} \\ 0 & \eta & 0 & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_1 \end{pmatrix},$$

$$D_9 = \begin{pmatrix} -\lambda_2 & 0 & 0 & 0 & 0 \\ -a_{15} & -a_{22} - a_{21} & a_{21} & a_{13} & -a_{14} \\ 0 & -\alpha & a_{21} - a_{11} & 0 & 0 \\ 0 & \eta & 0 & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_1 \end{pmatrix},$$

$$D_{10} = \begin{pmatrix} -\lambda_2 & 0 & 0 & 0 & 0 \\ -a_{15} & -a_{22} - a_{21} & a_{21} - \frac{\eta}{\alpha}a_{13} & a_{13} & -a_{14} \\ 0 & -\alpha & a_{21} - a_{11} & 0 & 0 \\ 0 & 0 & \frac{\eta}{\alpha}(a_{21} - a_{11} + \mu_2) & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_1 \end{pmatrix}.$$

Case 1. Let $\mu_2 \neq \mu_1$. Let $D_{11} = H_4D_{10}H_4^{-1}$, where H_4 is given in **step (1)**, and

$$D_{11} = \begin{pmatrix} -\lambda_2 & 0 & 0 & 0 & 0 \\ a_{15}\eta(a_{21} - a_{11} + \mu_2) & -a_1 & -a_2 & -a_3 & a_{14}\eta(a_{21} - a_{11} + \mu_2) \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_1 \end{pmatrix},$$

where a_1, a_2, a_3 are given in (4.3). Let $D_{12} = H_9D_{11}H_9^{-1}$, where

$$H_9 = \begin{pmatrix} a_{15}\eta(a_{21} - a_{11} + \mu_2) & -a_1 & -a_2 & -a_3 & a_{14}\eta(a_{21} - a_{11} + \mu_2) \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix},$$

$$D_{12} = \begin{pmatrix} -(a_1 + \lambda_2) & -(a_1\lambda_2 + a_2) & -(a_2\lambda_2 + a_3) & -a_3\lambda_2 & a_{14}\eta(a_{21} - a_{11} + \mu_2)(a_1 + \lambda_1) \\ 1 & 0 & 0 & 0 & a_{14}\eta(a_{21} - a_{11} + \mu_2) \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_1 \end{pmatrix}.$$

Then, by Lemma 4.1(2), for the matrix equation:

$$Q_0^2 + D_{12}\Sigma_{21} + \Sigma_{21}D_{12}^T = 0,$$

we can obtain the positive definite matrix as follows:

$$\Sigma_{21} = \begin{pmatrix} \frac{(a_1a_2 - a_3)\lambda_2^2 + a_2^2\lambda_2 + a_2a_3}{2H_{21}} & 0 & -\frac{a_2\lambda_2 + a_3}{2H_{21}} & 0 & 0 \\ 0 & \frac{a_2\lambda_2 + a_3}{2H_{21}} & 0 & -\frac{a_1 + \lambda_2}{2H_{21}} & 0 \\ -\frac{a_2\lambda_2 + a_3}{2H_{21}} & 0 & \frac{a_1 + \lambda_2}{2H_{21}} & 0 & 0 \\ 0 & -\frac{a_1 + \lambda_2}{2H_{21}} & 0 & \frac{a_1a_2 - a_3 + a_1\lambda_2(a_1 + \lambda_2)}{2a_3\lambda_2H_{21}} & 0 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix},$$

where $H_{21} = (a_1a_2 - a_3)(a_3 + \lambda_2a_2 + \lambda_2^2a_1 + \lambda_2^3)$.

Then, we have that $Q_2^2 + D\Sigma_2 + \Sigma_2D^T = 0$ is equivalent to

$$(H_8H_2H_3H_4H_9)Q_2^2(H_8H_2H_3H_4H_9)^T + D_{12}(H_8H_2H_3H_4H_9)\Sigma_2(H_8H_2H_3H_4H_9)^T + (H_8H_2H_3H_4H_9)\Sigma_2(H_8H_2H_3H_4H_9)^T D_{12}^T = 0,$$

which is further equivalent to

$$Q_0^2 + (a_{15}\eta(a_{21} - a_{11} + \mu_2)\sigma_2)^{-2}(D_{12}(H_8H_2H_3H_4H_9)\Sigma_2(H_8H_2H_3H_4H_9)^T + (H_8H_2H_3H_4H_9)\Sigma_2(H_8H_2H_3H_4H_9)^T D_{12}^T) = 0.$$

Therefore, we finally have that

$$\Sigma_2 = (a_{15}\eta(a_{21} - a_{11} + \mu_2)\sigma_2)^2(H_8H_2H_3H_4H_9)^{-1}\Sigma_{21}[(H_8H_2H_3H_4H_9)^{-1}]^T.$$

Thus, Σ_2 is calculated and is positive semi-definite.

Case 2. $\mu_1 = \mu_2$. Let $D_{13} = H_6D_{10}H_6^{-1}$, where H_6 is given in **step (1)**, and

$$D_{13} = \begin{pmatrix} -\lambda_2 & 0 & 0 & 0 & 0 \\ a_{15}\alpha & -a_{11} - a_{22} & -D^* & -a_{13}\alpha & a_{14}\alpha \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_1 \end{pmatrix}.$$

Let $D_{14} = H_{10}D_{13}H_{10}^{-1}$, where

$$H_{10} = \begin{pmatrix} a_{15}\alpha & -a_{11} - a_{22} & -D^* & -a_{13}\alpha & a_{14}\alpha \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{pmatrix}.$$

$$D_{14} = \begin{pmatrix} -(a_{11} + a_{22} + \lambda_2) & -((a_{11} + a_{22})\lambda_2 + D^*) & -D^*\lambda_2 & D_{71} & D_{73} \\ 1 & 0 & 0 & -a_{13}(\gamma + \alpha) & -a_{14}(\gamma + \alpha) \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & -\mu_2 & 0 \\ 0 & 0 & 0 & 0 & -\lambda_1 \end{pmatrix},$$

where $D_{73} = -a_{14}\alpha(a_{11} + a_{22} - \lambda_1)$. Then, by Lemma 4.1(1), for the matrix equation

$$Q_0^2 + D_{14}\Sigma_{22} + \Sigma_{22}D_{14}^T = 0,$$

we can obtain the positive definite matrix as follows:

$$\Sigma_{22} = \begin{pmatrix} \frac{(a_{11}+a_{22})\lambda_2+D^*}{2H_{22}} & 0 & -\frac{1}{2H_{22}} & 0 & 0 \\ 0 & \frac{1}{2H_{22}} & 0 & 0 & 0 \\ -\frac{1}{2H_{22}} & 0 & \frac{a_{11}+a_{22}+\lambda_2}{2D^*\lambda_2H_{22}} & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix}.$$

where $H_{22} = (a_{11} + a_{22} + \lambda_2)(a_{11} + a_{22})\lambda_2 + (a_{11} + a_{22})D^*$. Then, we have that $Q_2^2 + D\Sigma_2 + \Sigma_2D^T = 0$ is equivalent to

$$(H_8H_2H_3H_6H_{10})Q_2^2(H_8H_2H_3H_6H_{10})^T + D_{14}(H_8H_2H_3H_6H_{10})\Sigma_2(H_8H_2H_3H_6H_{10})^T + (H_8H_2H_3H_6H_{10})\Sigma_2(H_8H_2H_3H_6H_{10})^T D_{14}^T = 0,$$

which is further equivalent to

$$Q_0^2 + (a_{15}\alpha\sigma_2)^{-2}(D_{14}(H_8H_2H_3H_6H_{10})\Sigma_2(H_8H_2H_3H_6H_{10})^T + (H_8H_2H_3H_6H_{10})\Sigma_2(H_8H_2H_3H_6H_{10})^T D_{14}^T) = 0.$$

Therefore, we finally have that

$$\Sigma_2 = (a_{15}\alpha\sigma_2)^2(H_8H_2H_3H_6H_{10})^{-1}\Sigma_{22}[(H_8H_2H_3H_6H_{10})^{-1}]^T.$$

Thus, Σ_2 is calculated and is positive semi-definite. Then, $\Sigma = \Sigma_1 + \Sigma_2$ is positive definite. Therefore, the expression of a normal density function around the quasi-endemic equilibrium of model (1.2) is obtained by

$$\Phi(S, I, B, \gamma, c) = (2\pi)^{-\frac{5}{2}}|\Sigma|^{-\frac{1}{2}}e^{-\frac{1}{2}(S-S^*, I-I^*, B-B^*, \gamma-\gamma^*, c-c^*)\Sigma^{-1}(S-S^*, I-I^*, B-B^*, \gamma-\gamma^*, c-c^*)^T}.$$

This completes the proof. □

5. Numerical examples

In model (1.2), we take the parameters $A = 2, \beta = 0.5, K = 0.4, \mu_1 = 0.2, \mu_2 = 0.5, \alpha = 0.2, \eta = 0.2, \theta_1 = \theta_2 = 0.2, \bar{c} = 0.06, \bar{\gamma} = 0.1, b = 1$ and $\sigma_1 = \sigma_2 = 0.02$. The numerical simulation of solution $(S(t), I(t), B(t))$ is done with initial value $(S(0), I(0), B(0)) = (3, 1.5, 1.5)$. We obtain $R_0^S = 8.5436 > 1$. The positive equilibrium of model (1.2) is $E^* = (S^*, I^*, B^*, \gamma^*, c^*) \approx (4.08, 2.96, 1.18, 0.1, 0.06)$. (See Fig. 1) The conditions in Theorem 3.1 and Theorem 4.2 are satisfied. Therefore, there exists the stationary distribution of model (1.2) (See Fig. 2), and there is a unique normal density function near equilibrium E^* . Then, the positive definite matrix Σ is calculated as

$$\Sigma = \begin{pmatrix} 0.7302 & -0.4415 & -0.1544 & 0.1636 & 0.0413 \\ -0.4415 & 0.2971 & 0.0880 & -0.1090 & -0.0275 \\ -0.1544 & 0.0880 & 0.0352 & -0.0312 & -0.0079 \\ 0.1636 & -0.1090 & -0.0312 & 0.0500 & 0 \\ 0.0413 & -0.0275 & -0.0079 & 0 & 0.0500 \end{pmatrix},$$

and density function

$$\Phi(S, I, B, \gamma, c) = 1.0100 \times 10^{-6} e^{-\frac{1}{2}(S-4.08, I-2.96, B-1.18, \gamma-0.1, c-0.06) \times \Sigma^{-1} (S-4.08, I-2.96, B-1.18, \gamma-0.1, c-0.06)^T}$$

Then, S , I and B have the following marginal probability densities:

$$\frac{\partial \Phi}{\partial S} = 0.1863 e^{-49.4615(S-1.6875)^2}, \quad \frac{\partial \Phi}{\partial I} = 0.2920 e^{-28.5588(I-1.3961)^2}, \quad \frac{\partial \Phi}{\partial B} = 0.8483 e^{-408.5348(B-1.3961)^2}$$

The numerical simulation of the marginal probability densities for S , I and B is given in Figure 3.

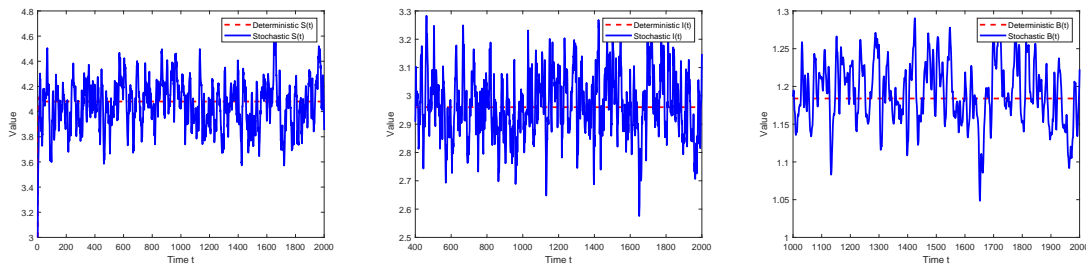


Figure 1. This shows the trajectories of $S(t)$, $I(t)$ and $B(t)$ of deterministic model (1.1) and stochastic model (1.2), respectively.

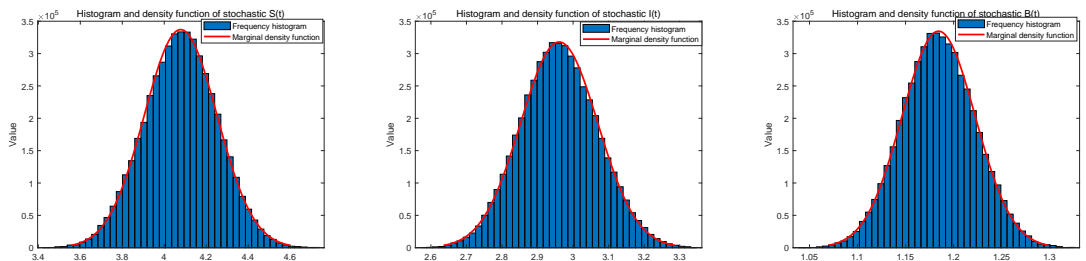


Figure 2. This shows that there exists the stationary distribution of model (1.2).

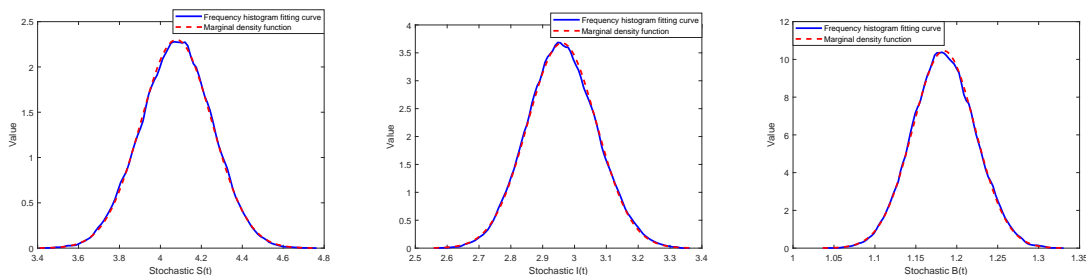


Figure 3. The numerical simulation of the marginal probability densities for S , I and B of model (1.2), respectively.

6. Conclusion

In this paper, we investigated a stochastic SIBS cholera model with saturation recovery rate and Ornstein-Uhlenbeck process. First, we proved that there is a unique global solution for any initial value of model. Secondly, the sufficient criterion of the stationary distribution of the model was obtained by constructing a suitable Lyapunov function, and the expression of the probability density function was calculated by the same condition. Finally, the correctness of the theoretical results is verified by numerical simulation, and the specific expression of the marginal probability density function is obtained.

This paper analyzes the interference of random disturbance in the process of recovering the infected person to the susceptible person during the transmission of cholera, but in fact, other parameters will also be more or less affected by random disturbance. In addition, people in many countries and regions have been vaccinated against cholera. In this process, cholera transmission is also affected by random factors. Therefore, there is still much work worthy of further study.

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Conflict of interest

The authors declare there is no conflict of interest.

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