



Research article

Forecasting volatility indices in stock and gold markets: Synergistic effects of the GARCH-MIDAS model and economic policy uncertainty

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Supplementary

Appendix A: Risk neutralization

We assume an exponential pricing kernel of the following form:

$$\begin{aligned}\zeta_{i+1,t} &= \frac{M_{i+1,t}}{E_{i,t}[M_{i+1,t}]} \equiv \frac{\exp(v_{i,t}\varepsilon_{i+1,t})}{E_t[\exp(v_{i,t}\varepsilon_{t,i+1})]} \\ &= \exp\left(v_{i,t}\varepsilon_{i+1,t} - \frac{1}{2}v_{i,t}^2\right)\end{aligned}$$

Under the no-arbitrage condition

$$E_{i,t}^Q[\exp(R_{i+1,t})] = E_{i,t}[\zeta_{i+1,t} \exp(R_{i+1,t})] = \exp(r)$$

Since

$$\begin{aligned}
E_{i,t}[\zeta_{i+1,t} \exp(R_{i+1,t})] &= E_{i,t} \left[\zeta_{i+1,t} \exp \left(r - \frac{1}{2} h_{i+1,t} + \lambda \sqrt{h_{i+1,t}} + z_{i+1,t} \right) \right] \\
&= E_{i,t} \left[\exp \left(v_{i,t} \varepsilon_{i+1,t} - \frac{1}{2} v_{i,t}^2 + r - \frac{1}{2} h_{i+1,t} + \lambda \sqrt{h_{i+1,t}} + \sqrt{h_{i+1,t}} \varepsilon_{i+1,t} \right) \right] \\
&= \exp \left(\frac{1}{2} (v_{i,t} + \sqrt{h_{i+1,t}})^2 - \frac{1}{2} v_{i,t}^2 + r - \frac{1}{2} h_{i+1,t} + \lambda \sqrt{h_{i+1,t}} \right) \\
&= \exp \left(v_{i,t} \sqrt{h_{i+1,t}} + \lambda \sqrt{h_{i+1,t}} + r \right)
\end{aligned}$$

This implies that

$$v_{i,t} \sqrt{h_{i+1,t}} + \lambda \sqrt{h_{i+1,t}} = 0$$

In order to determine the form of the risk-neutral distribution of the shocks, we consider the moment-generating function

$$\begin{aligned}
E_{i,t}^Q[\exp(u \varepsilon_{i+1,t})] \\
= \exp(u v_{i+1,t})
\end{aligned}$$

Under the Q measure, we have

$$\varepsilon_{i+1,t}^* = \varepsilon_{i+1,t} - v_{i+1,t}, \varepsilon_{i+1,t}^* \stackrel{iid^Q}{\sim} N(0,1)$$

Using the no-arbitrage condition above implies that

$$v_{i+1,t} = -\lambda$$

Now we can re-write the returns equation under the risk-neutral measure and condition variance as follows:

$$\begin{aligned}
R_{i+1,t} &= r - \frac{1}{2} h_{i+1,t} + \lambda \sqrt{h_{i+1,t}} + \varepsilon_{i+1,t} \sqrt{h_{i+1,t}} \\
&= r - \frac{1}{2} h_{i+1,t} + \sqrt{h_{i+1,t}} (\lambda + \varepsilon_{i+1,t}) \\
&= r - \frac{1}{2} h_{i+1,t} + \sqrt{h_{i+1,t}} \varepsilon_{i+1,t}^* \\
g_{i+1,t} &= \omega + \beta g_{i,t} + \alpha g_{i,t} \varepsilon_{i,t}^2 \\
&= \omega + \beta g_{i,t} + \alpha g_{i,t} (\varepsilon_{i,t}^* - \lambda)^2
\end{aligned}$$

Appendix B: VIX calculation under the risk-neutral GARCH-MIDAS models

Under the risk-neutral measure, we have:

$$\begin{aligned}
\frac{1}{252} \left(\frac{VIX_{i,t}^{Model}}{100} \right)^2 &\cong \frac{1}{22} \left[\sum_{k=1}^{M_t-i} E_{i,t}^Q [h_{i+k,t}] + \sum_{m=1}^{22-(M_t-i)} E_{i,t}^Q [h_{m,t+1}] \right] \\
&= \frac{1}{22} \left[\sum_{k=1}^{M_t-i} E_{i,t}^Q [\tau_t g_{i+k,t}] + \sum_{m=1}^{22-(M_t-i)} E_{i,t}^Q [\tau_{t+1} g_{m,t+1}] \right] \\
&= \frac{1}{22} \left[\sum_{k=0}^{M_t-1-i} \tau_t E_{i,t}^Q [g_{i+k+1,t}] + \sum_{m=0}^{21-(M_t-i)} \tau_{t+1} E_{i,t}^Q [g_{m+1,t+1}] \right]
\end{aligned}$$

Since

$$\begin{aligned}
E_{i,t}^Q [g_{i+k+1,t}] &= E_{i,t}^Q E_{i+k,t}^Q [g_{i+k+1,t}] \\
&= E_{i,t}^Q [\omega + \beta g_{i+k,t} + \alpha g_{i+k,t} (\varepsilon_{i+k,t}^* - \lambda)^2] \\
&= \Omega^* + \Gamma^* E_{i,t}^Q [g_{i+k,t}]
\end{aligned}$$

$$\begin{aligned}
E_{i,t}^Q [g_{m+1,t+1}] &= E_{i,t}^Q E_{m,t}^Q [g_{m+1,t+1}] \\
&= E_{i,t}^Q [\omega + \beta g_{m,t+1} + \alpha g_{m,t+1} (\varepsilon_{m,t+1}^* - \lambda)^2] \\
&= \Omega^* + \Gamma^* E_{i,t}^Q [g_{m,t+1}]
\end{aligned}$$

We can deduce that

$$\begin{aligned}
\frac{1}{22} \sum_{k=1}^{M_t-i} \tau_t E_{i,t}^Q [g_{i+k,t}] &= \frac{1}{22} \sum_{k=0}^{M_t-1-i} \tau_t E_{i,t}^Q [g_{i+k+1,t}] \\
&= \frac{1}{22} \sum_{k=0}^{M_t-1-i} \tau_t \{ \Omega^* + \Gamma^* E_{i,t}^Q [g_{i+k,t}] \} \\
&= \frac{1}{22} \sum_{k=0}^{M_t-1-i} \tau_t \{ \Omega^* + \Gamma^* [\Omega^* + \Gamma^* E_{i,t}^Q [g_{i+k-1,t}]] \} \\
&= \frac{1}{22} \sum_{k=0}^{M_t-1-i} \tau_t \{ \Omega^* (1 + \Gamma^*) + \Gamma^{*2} E_{i,t}^Q [g_{i+k-1,t}] \} \\
&= \frac{1}{22} \sum_{k=0}^{M_t-1-i} \tau_t \{ \Omega^* (1 + \Gamma^* + \dots + \Gamma^{*k-1}) + \Gamma^{*k} g_{i+1,t} \} \\
&= \frac{1}{22} \sum_{k=0}^{M_t-1-i} \tau_t \left\{ \frac{\Omega^* (1 - \Gamma^{*k})}{1 - \Gamma^*} + \Gamma^{*k} g_{i+1,t} \right\} \\
&= \tau_t \left[\frac{\Omega^*}{1 - \Gamma^*} \left(\frac{M_t - i}{22} - \frac{(1 - \Gamma^{*M_t-i})}{22(1 - \Gamma^*)} \right) + \frac{(1 - \Gamma^{*M_t-i})}{22(1 - \Gamma^*)} g_{i+1,t} \right] \\
&= \tau_t [A + B g_{i+1,t}]
\end{aligned}$$

$$\begin{aligned}
\frac{1}{22} \sum_{m=1}^{22-(M_t-i)} \tau_{t+1} E_{i,t}^Q[\mathbf{g}_{m,t+1}] &= \frac{1}{22} \sum_{m=0}^{21-M_t+i} \tau_{t+1} E_{i,t}^Q[\mathbf{g}_{m+1,t+1}] \\
&= \frac{1}{22} \sum_{m=0}^{21-M_t+i} \tau_{t+1} \left\{ \Omega^* + \Gamma^* E_{i,t}^Q[\mathbf{g}_{m,t+1}] \right\} \\
&= \frac{1}{22} \sum_{m=0}^{21-M_t+i} \tau_{t+1} \left\{ \Omega^* + \Gamma^* [\Omega^* + \Gamma^* E_{i,t}^Q[\mathbf{g}_{m-1,t+1}]] \right\} \\
&= \frac{1}{22} \sum_{m=0}^{21-M_t+i} \tau_{t+1} \left\{ \Omega^* (1 + \Gamma^*) + \Gamma^{*2} E_{i,t}^Q[\mathbf{g}_{m-1,t+1}] \right\} \\
&= \frac{1}{22} \sum_{m=0}^{21-M_t+i} \tau_{t+1} \left\{ \Omega^* (1 + \Gamma^* + \dots + \Gamma^{*M_t-1-i+m}) + \Gamma^{*M_t-i+m} \mathbf{g}_{i+1,t} \right\} \\
&= \frac{1}{22} \sum_{m=0}^{21-M_t+i} \tau_{t+1} \left\{ \Omega^* \frac{(1 - \Gamma^{*M_t-i+m})}{(1 - \Gamma^*)} + \Gamma^{*M_t-i+m} \mathbf{g}_{i+1,t} \right\} \\
&= \tau_{t+1} \left[\frac{\Omega^*}{1 - \Gamma^*} \left(\frac{22 - M_t + i}{22} - \Gamma^{*M_t-i} \frac{(1 - \Gamma^{*22-M_t+i})}{22(1 - \Gamma^*)} \right) + \Gamma^{*M_t-i} \frac{(1 - \Gamma^{*22-M_t+i})}{22(1 - \Gamma^*)} \mathbf{g}_{i+1,t} \right] \\
&= \tau_{t+1} [C + D \mathbf{g}_{i+1,t}]
\end{aligned}$$

Finally, we get the expression for VIX as follows:

$$\begin{aligned}
\frac{1}{252} \left(\frac{VIX_{i,t}^{Model}}{100} \right)^2 &\cong \frac{1}{22} \left[\sum_{k=1}^{M_t-i} E_{i,i}^Q[h_{i+k,t}] + \sum_{m=1}^{22-(M_t-i)} E_{i,i}^Q[h_{m,t+1}] \right] \\
&= \tau_t [A + B \mathbf{g}_{i+1,t}] + \tau_{t+1} [C + D \mathbf{g}_{i+1,t}]
\end{aligned}$$

where M_t denotes the number of days in the t month, A , B , C , and D are functions of the

parameters, satisfying $A = \frac{\Omega^*}{1 - \Gamma^*} \left(\frac{M_t - i}{22} - B \right)$, $B = \frac{1 - \Gamma^{*M_t-i}}{22(1 - \Gamma^*)}$,

$C = \frac{\Omega^*}{1 - \Gamma^*} \left(\frac{22 - M_t + i}{22} - D \right)$, and $D = \Gamma^{*M_t-i} \frac{1 - \Gamma^{*22-M_t+i}}{22(1 - \Gamma^*)}$, $\Omega^* > 0, \Omega^* = \omega$. The persistence of the

conditional variance satisfies $\Gamma^* = \beta + \alpha(1 + \lambda^2) < 1$.



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